

## **Aerobic fitness thresholds to define poor cardiometabolic health in children and youth**

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## Abstract

Aerobic fitness is an apparent candidate for screening children and youth for poor cardiometabolic health and future risk of cardiovascular disease (CVD). Yet, age- and sex-specific cut points for children and youth determined using a maximal protocol and directly measured peak oxygen consumption ( $VO_{2peak}$ ) does not exist. We used a nationally representative sample of 1462 Norwegian children and youth (788 boys and 674 girls aged 8.7–10.4 years and 14.7–16.7 years) who in 2005–2006 performed a maximal cycle ergometer test with direct measurement of  $VO_{2peak}$ , along with measurement of several other risk factors for CVD (systolic blood pressure, waist circumference:height ratio, total:high density lipoprotein cholesterol ratio, triglycerides, Homeostasis Model Assessment for Insulin Resistance). Based on the proportion of children having clustering (least favorable quartile) of 6 (1.6%),  $\geq 5$  (5.2%), and  $\geq 4$  (10.6%) CVD risk factors, we established the 2<sup>nd</sup>, 5<sup>th</sup>, and 10<sup>th</sup> percentile cut points for  $VO_{2peak}$  (ml/kg/min) for children and youth aged 8–18 years. Classification accuracy was determined using the Kappa coefficient ( $k$ ), sensitivity and specificity. For boys, the 2<sup>nd</sup>, 5<sup>th</sup>, and 10<sup>th</sup> percentile  $VO_{2peak}$  cut points were 33.6–36.4, 36.3–39.8, and 38.7–43.0 ml/kg/min, respectively. For girls, the corresponding cut points were 29.7–29.1, 32.4–31.4, and 34.8–33.5 ml/kg/min, respectively. Together with BMI, but without more invasive measures of traditional risk factors for CVD, these cut points can be used to screen schoolchildren for poor cardiometabolic health with moderate discriminating ability ( $k \leq 0.53$ ).

**Keywords:** maximal oxygen consumption; cardiovascular disease; schoolchildren; screening; prevention

## Introduction

There is consistent evidence that increased aerobic fitness relates to better cardiometabolic health in children and youth<sup>1-3</sup>, and that this relation persists into adulthood<sup>4,5</sup>. Consequently, aerobic fitness is a candidate for screening and early intervention to prevent future metabolic disturbances and cardiovascular disease (CVD). Aerobic fitness<sup>6</sup>, especially in combination with fatness<sup>7,8</sup>, can be used to identify at-risk individuals, without requiring invasive measures of traditional CVD risk factors (blood samples). However, for measures of aerobic fitness to be used as an indicator of cardiometabolic risk and be decisive in terms of initiating preventive initiatives to increase physical activity, cut points to define children with poor cardiometabolic health are needed.

Ruiz et al<sup>6</sup> reviewed the evidence for  $VO_{2peak}$  (ml/kg/min) cut points established to indicate metabolic risk in children and youth aged 8 to 19 years in 2016. Based on seven included studies, they concluded that  $VO_{2peak}$  levels lower than 41.8 to 47.0 ml/kg/min in boys and lower than 34.6 to 39.5 ml/kg/min in girls indicated poor cardiometabolic health. Yet, this study has two major limitations. First, all seven studies included in the systematic review estimated  $VO_{2peak}$  based on indirect performance measures, some known to underestimate  $VO_{2peak}$  and some having unknown validity. For example, Mesa et al<sup>9</sup>, Moreira et al<sup>10</sup>, and Ruiz et al<sup>11</sup> used the 20 meter multistage shuttle run test (MSRT) to estimate  $VO_{2peak}$  applying the Leger<sup>12</sup> and Matsuzaka<sup>13</sup> equations, which are both shown to underestimate  $VO_{2peak}$  in children and adolescents<sup>14,15</sup>. Furthermore, we are not aware of any cross-validation studies of the submaximal treadmill protocol used to estimate  $VO_{2peak}$  in the studies based on the US National Health and Nutrition Examination Survey (NHANES)<sup>16,17</sup>. Second, the review by Ruiz et al<sup>6</sup> did not account for age, which imposes a clear limitation as it is well-known that  $VO_{2peak}$  per kg body mass increases with age during adolescence in boys, whereas it slightly decreases in girls<sup>18</sup>. Thus, accounting for age is critical to attain useful cut points. This is supported by the findings by Adegboye et al<sup>19</sup>, the largest study included in the systematic review by Ruiz et al<sup>6</sup>, showing that  $VO_{2peak}$  cut points indicating poor cardiometabolic health increased with age in boys (43.6 ml/kg/min in 9-year-olds and 46.0 ml/kg/min in 15-year-olds), but decreased with age in girls (37.4 ml/kg/min in 9-year-olds and 33.0 ml/kg/min in 15-year-olds). Similar findings were shown by Welk et al<sup>17</sup>, which is the only study to have established cut points by sex and age between ages 10 and 18 years. These cut points have been adapted by FITNESSGRAM<sup>20</sup> but a limitation of the study by Welk et al<sup>17</sup> is reliance on a submaximal test of  $VO_{2peak}$ .

To our knowledge, aerobic fitness cut points established using a large sample of children performing maximal tests with direct measurement of  $VO_{2peak}$  does not exist besides the study by Adegboye et al<sup>19</sup>, which included directly measured  $VO_{2peak}$  from the *Physical Activity among Norwegian Children*

*Study* (PANCS)<sup>18 21</sup>, as a part of their study sample. However, Adegboye et al<sup>19</sup> did not present age-specific cut points beyond the age groups of 9- and 15-year-olds. Thus, the aim of this paper was to define age- and sex-specific cut points for  $VO_{2peak}$  (ml/kg/min) indicative of poor cardiometabolic health in children and youth aged 8 to 18 years old solely from the PANCS study.

## **Methods**

### **Participants**

We included a nationally representative sample of Norwegian children and adolescents aged 8.7 to 10.4 years (“9-year-olds”) and 14.7 to 16.7 years (“15-year-olds”) from PANCS<sup>18 21</sup>, conducted during 2005–2006. Statistics Norway selected the cohort by cluster sampling with schools as the primary unit. When a school agreed to participate, we invited all children in grade 4 and grade 10. Thus, we invited a total of 2818 children and adolescents, of whom 2299 accepted and participated in the study. Of these, 1462 children (788 boys and 674 girls) provided data for the present analyses.

Procedures and methods conform to ethical guidelines defined by the World Medical Association’s Declaration of Helsinki and its subsequent revisions. The Regional Committee for Medical and Health Research Ethics and Norwegian Social Sciences Data Services approved the study. We obtained written informed consent from each participant’s parents or legal guardian prior to testing.

### **Procedures**

We have previously published detailed descriptions of the PANCS studies<sup>18 21</sup> and therefore provide only a brief overview of the relevant procedures herein.

$VO_{2peak}$  (ml/min/kg) was directly assessed using a portable MetaMax III X oxygen analyzer (Cortex Biophysics, Leipzig, Germany) during a progressive cycle test until volitional exhaustion using an electronically braked cycle ergometer (Ergonomic 839E; Monark, Varberg, Sweden). Initial and incremental work rates were 20 Watt (W) for 9-year-olds weighing <30 kg, 25 W for 9-year-olds weighing  $\geq$ 30 kg, 40 W for 15-year-old girls and 50 W for 15-year-old boys. Work rate increased every third minute until exhaustion. We recorded heart rate throughout the test using a heart rate monitor (Polar Vantage, Finland) and  $VO_2$ , respiratory exchange ratio, and ventilation every 10 seconds during the last minutes of the test. We defined  $VO_{2peak}$  as the mean of the three highest consecutive measurements. Every morning we calibrated the analyzer against known gas mixtures, and barometric pressure against values from the local weather station. The cycle ergometer was

electronically calibrated every morning and mechanically calibrated after being moved. The primary criterion of an acceptable test (maximal effort) was that the participants demonstrated clear signs of intense effort and clear symptoms of fatigue (e.g., facial flushing or difficulties in maintaining pedaling frequency). In addition, attainment of one objective criterion (heart rate  $\geq 185$  beats/minute or a respiratory exchange ratio  $\geq 0.99$ ) was needed to accept the test as valid. Body mass and height was measured using standardized procedures (Seca 770, SECA GmbH, Hamburg, Germany) with children wearing light clothing. Body mass index (BMI) ( $\text{kg}\cdot\text{m}^{-2}$ ) was calculated, and individuals' BMI statuses classified according to the BMI criteria by Cole et al <sup>22</sup>. Waist circumference (WC) was measured with a metal anthropometric tape, taken midway between the lower rib and the iliac crest at the end of a normal expiration. After being seated for a minimum of 5 minutes, we measured systolic blood pressure (SBP) five times at 2-minute intervals using an automated blood pressure monitor (Omega™ Noninvasive blood pressure monitor; Invivo research, Inc., Orlando, FL). We used the mean of the last three measurements for analysis. Following an overnight fast, venous blood samples were collected between 08:00 and 10:00 a.m. The samples were analyzed for total cholesterol (TC), triglycerides (TG), high-density lipoprotein (HDL) cholesterol, and glucose at the Central Laboratory of Ullevaal University Hospital (Oslo, Norway) using Routine enzymatic colorimetric assays from Roche Diagnostics and performed on a Cobas Integra analyser (F. Hoffmann-La Roche Ltd, Basel, Switzerland). Insulin was analyzed at the Aker University Hospital (Oslo, Norway) by fluoroimmunoassay using an automatic immunoassay system (AutoDELFIAHInsulin; PerkinElmer, Turku, Finland). We calculated the TC:HDL ratio and Homeostasis Model Assessment for Insulin Resistance (HOMA) ( $\text{glucose (mmol/L)} * \text{insulin (pmol/L)} / 22.5$ ) <sup>23</sup>.

### Statistical analyses

Participants' characteristics were reported as frequencies, means and standard deviations (SD), or medians and interquartile ranges (IQR) if variables were skewed. We tested for differences in characteristics between included and excluded individuals using an independent samples t-test (normally distributed variables) or a Mann-Whitney test (skewed variables), after correcting for age and sex (i.e., using residuals from a linear regression model).

Similar to a previous study <sup>8</sup>, we used clustering of CVD risk factors as the criterion measure for defining children with poor cardiometabolic health. Clustering was defined as present in an individual if their measurements were within the least favourable quartile (at risk) for a greater than expected number of the six risk factors: SBP, TG, TC:HDL ratio, HOMA, WC:height ratio, and  $\text{VO}_{2\text{peak}}$ , given independence of the risk factors. The observed proportion that had clustering was compared to the

expected proportion given independence among the risk factors using the binomial formula ( $\frac{n! * p^r * (1 - p)^{n-r}}{r! * (n - r)!}$ ), where  $n$  is the possible number of risk factors (6),  $p$  is the probability of having a risk factor (0.25), and  $r$  is the number of the risk factors for which the probability is calculated (0 through 6). The expected proportions having 0 to 6 risk factors were 0.1780, 0.3560, 0.2966, 0.1318, 0.0330, 0.0044, and 0.0002, respectively. We then calculated the Odds Ratios (OR) and 95% confidence intervals (CI) for the observed vs. expected proportions having clustering of 3, 4, 5, and 6 risk factors as  $\exp(\ln[OR] \pm 1.96 * SE[\ln(OR)])$ , where  $SE[\ln(OR)] = \sqrt{[1/n_i] + [1/N - n_i]}$ , and  $N$  is the total number of children and  $n_i$  is the number of children with the specific number of risk factors.

We used linear regression to determine the relation between age and  $VO_{2peak}$  for boys and girls separately as this relation was gender-specific (sex\*age interaction:  $p < .001$ ). Based on these equations we calculated the mean  $VO_{2peak}$  values for children aged 8 to 18. As the SD increased with age in boys and decreased with age in girls, we calculated age- and sex-specific SDs by establishing the slope with age in boys and girls. We thereafter calculated the cut points for  $VO_{2peak}$  being below the 2<sup>nd</sup>, 5<sup>th</sup>, and 10<sup>th</sup> percentile using one-tailed z-scores of -2.05, -1.65, and -1.28, respectively. These proportions correspond to participants defined as at-risk having 6 (1.6%),  $\geq 5$  (5.2%), and  $\geq 4$  (10.6%) risk factors (upper quartile), for which observed clustering of risk factors was significantly higher than expected.

The probability of being at risk using clustering (upper quartile) of  $\geq 4$ ,  $\geq 5$ , or 6 risk factors as the criterion measure was estimated by logistic regression (OR, 95% CI) for 3 different classifications based on  $VO_{2peak}$  ( $< 10^{th}$ ,  $< 5^{th}$ , and  $< 2^{nd}$  percentile for  $VO_{2peak}$ ), 2 different classifications based on BMI (overweight or obese, and obese), and the 6 combined classifications (3 for  $VO_{2peak}$  \* 2 for BMI). Classification accuracy was reported as sensitivity (the true positive/total positive rate), specificity (the true negative/total negative rate), and Kohen's kappa ( $k$ ). Kohen's kappa was interpreted as fair for  $k = 0.21-0.40$  and moderate for  $k = 0.41-0.60$  <sup>24</sup>.

Analyses were performed using IBM SPSS v. 23 (IBM Corporation, Software Group, Somers, New York, USA). A p-value  $< .05$  indicated statistically significant findings.

## Results

### Participants' characteristics

Of the 1306 9-year-olds and 993 15-year-olds included in the PANCS study, 843 (65%) 9-year-olds and 619 (62%) 15-year-olds provided valid data on all relevant variables and were included in the present analyses (Table 1). Individuals included (n = 1462) and excluded from the analyses (n = 837; n = 291–837 for different variables in the attritional analyses) was similar with respect to proportion of boys and girls, age, body mass, BMI, WC, SBP, and parents' level of education (p >.170). The proportion of individuals with Caucasian origin was higher among those included (91 vs. 88%, p = .045). Height (mean (SD) 153.1 (17.5) vs. 152.2 (17.3) cm, p < .001), WC:height ratio (mean (SD) 0.441 (0.047) vs. 0.447 (0.052), p = .007), TC:HDL ratio (median (IQR) 2.54 (0.76) vs. 2.62 (0.83), p < .001), TG (median (IQR) 0.63 (0.34) vs. 0.67 (0.39) mmol/l, p < .001), and HOMA (median (IQR) 7.86 (6.54) vs. 8.63 (8.65), p = .003) differed significantly in favour of the included children. Absolute VO<sub>2peak</sub> was higher in the excluded individuals (median (IQR) 1.81 (1.17) vs. 1.78 (1.19) l/min, p = .004), whereas VO<sub>2peak</sub> relative to body mass tended to be higher in the included individuals (mean (SD) 46.5 (8.08) vs. 45.9 (8.2) ml/kg/min, p = .064).

The observed number of individuals with clustering of risk factors was higher than expected for 4 or more risk factors. 71 (4.9%) individuals had clustering of 4 risk factors (OR = 1.48, 95% CI 1.17–1.88), 49 (3.4%) individuals had clustering of 5 risk factors (OR = 7.73, 95% CI 5.81–10.27), and 23 (1.6%) individuals had clustering of 6 risk factors (OR = 80.0, 95% CI 53.0–121). Thus, we established VO<sub>2peak</sub> cut points for the proportion of individuals having clustering of ≥ 4 risk factors (n = 143, 10.6%, OR = 2.82, 95% CI 2.37–3.35), ≥ 5 risk factors (n = 75, 5.2%, OR = 11.30, 95% CI 8.92–14.33), and 6 risk factors (n = 23, 1.6%, OR = 80.0, 95% CI 53.0–121). Proportions were similar across sex and age groups (p = .579, Table 1).

#### Cut points for VO<sub>2peak</sub>

Based on the proportions of individuals with clustering of ≥ 4 (10.6%), ≥ 5 (5.2%), and 6 (1.6%) risk factors, we established the 2<sup>nd</sup>, 5<sup>th</sup>, and 10<sup>th</sup> percentile cut points for VO<sub>2peak</sub> using the respective one-tailed z-scores (Table 2). VO<sub>2peak</sub> increased by 0.66 ml/kg/min per year for boys (VO<sub>2peak</sub> = 41.99806 + 0.66173\*age, R<sup>2</sup> = 0.066, p < .001) and decreased by 0.25 ml/kg/min per year (VO<sub>2peak</sub> = 45.32124 – 0.25030\*age, R<sup>2</sup> = .014, p = .002) for girls. The mean SD was 7.46 and increased by 0.19 ml/kg/min per year in boys, whereas the mean SD was 6.26 and decreased by 0.09 ml/kg/min per year in girls. The resulting cut points for VO<sub>2peak</sub> increased with age for boys (2<sup>nd</sup> percentile: 0.28 ml/kg/min per year; 5<sup>th</sup> percentile: 0.35 ml/kg/min per year; 10<sup>th</sup> percentile: 0.42 ml/kg/min per year), but decreased with age for girls (2<sup>nd</sup> percentile: -0.06 ml/kg/min per year; 5<sup>th</sup> percentile: -0.10 ml/kg/min per year; 10<sup>th</sup> percentile: -0.13 ml/kg/min per year).

Probability and classification accuracy of cardiometabolic risk for different VO<sub>2peak</sub> and BMI indices

Table 3 shows the probability of being classified with clustering of risk factors along with sensitivity, specificity, and the Kappa coefficient based on different VO<sub>2peak</sub> and BMI indices. Overall, 1.6–10.6% of the individuals were defined as at risk based on actual observed clustering of ≥ 4–6 CVD risk factors, whereas the corresponding proportions defined as at risk using the 11 VO<sub>2peak</sub> and BMI indices of cardiometabolic risk varied from 1.0 to 14.2%. All indices of cardiometabolic risk significantly increased the odds of poor metabolic health as defined by any outcome. The probability of being classified as at risk based on VO<sub>2peak</sub> varied from OR 14.2 to 33.0 among the outcome criteria. Overweight or obesity (ORs 21.3–44.6), and especially obesity (63.2–91.7), showed stronger associations with metabolic risk than VO<sub>2peak</sub> for all outcomes. The combined VO<sub>2peak</sub> and BMI indices increased the odds of cardiometabolic risk beyond the odds of the respective variables separately for some indices, but this pattern was not consistent. Low VO<sub>2peak</sub> AND overweight or obesity resulted in ORs 20.9–66.3, whereas low VO<sub>2peak</sub> AND obesity resulted in ORs 21.7–153 across outcomes.

Classification accuracy varied widely across all outcomes and VO<sub>2peak</sub> and BMI indices ( $k = 0.15–0.53$ ). Among the VO<sub>2peak</sub> indices, the 10<sup>th</sup> percentile was the best index for determining clustering of ≥ 4 risk factors ( $k = 0.41$ ), the 5<sup>th</sup> and 10<sup>th</sup> percentile were similar for determining clustering of ≥ 5 risk factors ( $k = 0.38–0.39$ ), and the 2<sup>nd</sup> percentile was the best index for determining clustering of 6 risk factors ( $k = 0.22$ ). Among the BMI indices, overweight or obese was the best index for determining clustering of ≥ 4 risk factors ( $k = 0.48$ ), whereas obesity performed best for clustering of ≥ 5 risk factors ( $k = 0.49$ ) and 6 risk factors ( $k = 0.39$ ). The highest kappa was found for VO<sub>2peak</sub> below the 10<sup>th</sup> percentile AND overweight or obese with clustering of ≥ 4 ( $k = 0.48$ ) and ≥ 5 ( $k = 0.53$ ) risk factors. Kappa coefficients were ≤ 0.39 for all indices for clustering of 6 risk factors. Sensitivity and specificity varied with the proportion classified at risk by the indices and outcomes.

## Discussion

We are the first to present age- and sex-specific cut points for VO<sub>2peak</sub> (ml/kg/min) to define children and youth with poor cardiometabolic health based on a large sample and using a maximal test with direct measurement of oxygen consumption. These cut points, together with BMI, can be used to screen children for poor cardiometabolic health.



Ruiz et al<sup>6</sup> recently summarized previous studies that have sought to determine  $VO_{2peak}$  cut points to indicate cardiometabolic risk in children and youth. However, these results are of limited usefulness because they did not account for the age-related development of  $VO_{2peak}$  per kg body mass<sup>18</sup>, which is a key feature during childhood growth and development. Moreover, there are at least two major challenges and inconsistencies when constructing and comparing cut points across previous studies. One relates to the operationalization or definition of cardiometabolic risk or the metabolic syndrome (who should be targeted for intervention?), and the other relates to the proportion of children actually detected by screening, which most often correspond poorly with the target population in previous studies (who is targeted for intervention through the screening?).

Our findings show that classification accuracy for most  $VO_{2peak}$  and BMI indices were fair to moderate. Yet, it should be mentioned that the Kappa statistic provides lower values with higher class skew between positive and negative cases<sup>24</sup>, which is a challenge in the present analysis where small proportions were defined as at risk across the criterion measures (1.6–10.6%). Nevertheless, sensitivity (the ability to flag those at risk) increased and specificity (the ability to exclude those not at risk) decreased when classifying a greater proportion as at risk. This is a key challenge when screening for risk/disease, and thus the crucial question arises: what proportion of children and youth should we classify as at risk? As there is no universally accepted standard for defining metabolic syndrome or poor cardiometabolic health in children and youth, the proportion of children defined as at risk varies widely between studies. Welk et al<sup>17</sup> defined 6% as at risk based on having  $\geq 3$  of 5 risk factors (WC, blood pressure, HDL, TG, glucose) using the National Cholesterol Education Program/Adult Treatment PANEL III criteria. Ruiz et al<sup>11</sup> defined 15% as at risk based on achieving ideal levels for  $\geq 4$  of 7 health behaviors/factors (smoking, BMI, physical activity, diet, glucose, blood pressure, and TG) using cut points suggested by the American Heart Association. Using pooled data from 15794 6–18-year-olds, Andersen et al<sup>16</sup> showed that the prevalence of metabolic syndrome as defined by the IDF criteria was less than 1%, whereas 6.2 and 15.7% were classified as at risk of poor cardiometabolic health when defined as exhibiting clustering (upper quartile) of  $\geq 4$  and  $\geq 3$  of 5 risk factors (WC, SBP, inverse HDL, TG, and HOMA), respectively. A common approach to avoid disagreement when applying absolute cut points for CVD risk factors in children and youth has been to define individuals scoring  $< -1$  SD (16%) on a continuous composite score (sum of z-scores) as at risk<sup>10 16 19 25</sup>. We used clustering of 6 (1.6% of individuals),  $\geq 5$  (5.2% of children), and  $\geq 4$  (10.6% of children) of 6 risk factors (WC:height ratio, SBP, TC:HDL ratio, TG, HOMA and inverse  $VO_{2peak}$ ) as a basis for defining the 2<sup>nd</sup>, 5<sup>th</sup>, and 10<sup>th</sup> percentile  $VO_{2peak}$  cut points. Thus, similar to Andersen et al<sup>8</sup>, we used a biological basis for defining at-risk individuals, by determining the proportions exhibiting clustering of more risk factors than expected given independence among risk factors. Nonetheless, to

facilitate comparability among studies we have provided age- and sex-specific mean  $VO_{2peak}$  values along with SDs, allowing for calculation of any percentile in the present sample.

A further challenge in the existing literature is that most previous studies have applied Receiver Operating Characteristics (ROC) curve analyses to determine the best cut points for a healthy  $VO_{2peak}$ . By definition, when using the ROC approach, the best cut point is selected as the one that balances sensitivity and specificity. Due to the class skew (e.g., 10 % at risk; 90% not at risk) and the moderate relationship between the exposure and the outcome, achieving a balanced true positive and true negative rate leads to a “right-skew” of the cut points, resulting in cut points that define an exaggerated proportion of the sample at risk (Table 4), which leads to a high absolute false positive rate. Thus, although most previous studies have defined metabolic risk as scoring below  $< -1$  SD on a continuous composite score, by definition classifying 16% of individuals at risk, the cut points established have selected 9–22<sup>19</sup>, 21–37<sup>16</sup>, and 44–49%<sup>25</sup> as at risk. Similarly, Welk et al<sup>17</sup> classified 6.3 and 5.9% of boys and girls with the metabolic syndrome, whereas their cut points classified 33–48% as at risk. Also, Ruiz et al<sup>11</sup> classified 13 and 16% of boys and girls with cardiometabolic risk, whereas their cut points classified 30–38% as at risk. As can be seen in Table 4, the cut points suggested by Ruiz et al<sup>25</sup> and Welk et al (“healthy fitness zone”)<sup>17</sup> are close to the mean values of the samples, also found by Mesa et al<sup>9</sup>. This implies that most previous cut points are substantially overestimated (too high) in terms of identifying the target population defined to have poor cardiometabolic health. The lower proportion classified as at risk explains the lower sensitivity and higher specificity for most  $VO_{2peak}$  and BMI indices in the present study, compared to these previous studies<sup>9 11 16 17 25</sup>.

Discriminating between children that are at risk and children not at risk is a challenging task given the trade-off between sensitivity and specificity, which consequently affects the proportion of children that could be targeted for preventive initiatives. The combined  $VO_{2peak}$  and BMI indices provided the best classification accuracy, with estimates similar to a previous study investigating the performance of a noninvasive composite score using the Andersen aerobic intermittent running field test and BMI in 10-year-old children ( $k = 0.52–0.53$ )<sup>7</sup>. In a class of 30 schoolchildren, assuming 10% are at risk as defined by having clustering of  $\geq 4$  of 6 CVD risk factors in the present study, defining children as at risk according to the index with the highest kappa ( $k = 0.48$ ), the  $VO_{2peak}$  10<sup>th</sup> percentile AND *overweight/obese* index (6.0% of children at classified at risk; sensitivity 61%; specificity 97%) could classify ( $\approx$ )2 out of 3 children correctly (true positive) and 1 out of 27 children incorrectly (false positive) as at risk. These numbers would obviously change with a stricter or more liberal cut point. Ultimately, the choice of  $VO_{2peak}$  cut points (and cardiometabolic risk criteria) must be a

comprehensive consideration of the pros and cons of misclassification, including the health risk of being erroneously missed (false negative), the worry and/or stigmatization of being erroneously detected (false positive), and the type of preventive action that can be offered.

Of great importance for screening,  $VO_{2peak}$  can be estimated from different field tests, which makes aerobic fitness both a simple and feasible indicator of cardiometabolic risk, especially in a school setting where more invasive measures of traditional CVD risk factors are less practical to obtain. However, estimation of  $VO_{2peak}$  from field tests is limited by large prediction errors on the individual level<sup>14 15 26</sup>. Therefore, future studies should determine cut points for performance measures directly using for example the MSRT (laps or speed)<sup>12</sup> or the Andersen test (meters)<sup>27</sup>. In fact, using these tests can improve the prediction of metabolic health in children, as it has been shown that the Andersen test associates more strongly with metabolic health than does  $VO_{2peak}$ <sup>28</sup>. Derived from our findings that 10.6 and 5.2% of children across age had clustering of  $\geq 4$  and 5 risk factors (which were thus used as the criteria for our  $VO_{2peak}$  cut points), we suggest the 5<sup>th</sup> and 10<sup>th</sup> percentile thresholds as determined from the international MSRT references suggested by Tomkinson et al<sup>29</sup>, could indicate increased metabolic risk. Yet, these proportions differ substantially from those suggested by Tomkinson et al, derived from the FITNESSGRAM cut points<sup>17 20</sup>, for reasons discussed above.

## Perspectives

Given the indisputable health effects of regular physical activity<sup>30 31</sup>, one might argue increasing sensitivity at the expense of decreasing specificity by choosing, for example, the 50<sup>th</sup>  $VO_{2peak}$  percentile cut point as indicative of CVD risk, would be a minor problem. The 50<sup>th</sup> percentile in the present study led to sensitivity of 93.0, 97.2, and 100% (specificity 54.9, 52.7, and 51.0%), respectively, using clustering of  $\geq 4$ ,  $\geq 5$ , and 6 CVD risk factors as the outcomes (results not shown). However, targeting half the population with physical activity initiatives to capture most individuals at risk points toward a population strategy rather than a high-risk strategy for prevention<sup>32</sup>. As such, if using a population strategy that targets all individuals, screening would be of minor relevance. Both strategies have their advantages and disadvantages, and both are likely needed in an effort to prevent poor health<sup>32</sup>. Importantly, increased school-based physical activity can significantly increase aerobic fitness and improve cardiometabolic health<sup>33-35</sup>, with greater effects in the children having the least favorable cardiometabolic profile<sup>33</sup>. This finding demonstrates that high-risk children can be reached through a school-based population approach to combat low physical activity levels among children and youth<sup>36</sup>.

## Strengths and limitations

This study has several strengths. Most important is the use of a large, nationally representative sample of children and youth performing a maximal test with direct measurement of oxygen consumption to determine  $VO_{2peak}$ . The maximal effort is verified by high peak heart rates and respiratory exchange ratio values across the age and sex groups. Yet, we used a cycle ergometer test, which is known to produce lower values than uphill treadmill running in both children (5% lower)<sup>37</sup> and adults (7% lower)<sup>38,39</sup>. Thus, raising our cut points  $\approx 2\text{--}3$  ml/kg/min would make them equivalent to values obtained by a treadmill protocol. Further, while most previous studies<sup>6,9-11,16</sup>, including the systematic review by Ruiz et al<sup>6</sup>, only suggest cut points for age groups, the age- and sex-specific cut points provided herein allow for a more precise and accurate identification of risk at all ages.

Moreover, researchers and clinicians can use the present findings to classify schoolchildren of all ages as at risk according to the 2<sup>nd</sup>, 5<sup>th</sup>, or 10<sup>th</sup> percentile, or alternatively, by using the means and SDs reported to calculate any given percentile according to  $VO_{2peak}$  based on this representative sample. However, it should be borne in mind that the age range included in this study comprised children aged 8.7 to 10.4 years and 14.7 to 16.7 years, and that trends by age could be non-linear. Additionally, on an individual level, maturation and pubertal development could influence both CVD risk factors and  $VO_{2peak}$  beyond chronological age. Yet, we did not account for these factors because age and maturation status (analyzed by Tanner stage and peak height velocity offset) was completely collinear ( $r \geq 0.97$ ) in the current study, due to the study sampling. Moreover, applying this information would reduce the clinical value of the current data, for example if used for screening in a school setting, as this information might not be readily available.

A limitation of our study is the cross-sectional design, which precludes determination of the predictive validity of the suggested  $VO_{2peak}$  cut points. Thus, future studies should apply and cross-validate the cut points in new samples and using longitudinal study designs. Furthermore, how to verify a maximal (peak) effort on graded exercise tests in children is a matter of debate<sup>40</sup>. Our secondary criteria for accepting a  $VO_{2peak}$  test may seem rather low. However, due to the large individual variation in maximal (peak) values, we used these criteria to avoid excluding participants erroneously. Moreover, our peak heart rate levels, obtained from cycle ergometry, are comparable to other studies in children using both treadmill and cycle ergometry (200-204 beats/min)<sup>41-43</sup>, despite cycle ergometry being known to provide lower maximal values than treadmill exercise<sup>40</sup>.

Finally, we did not measure the participants' lean body mass. Reporting  $VO_{2peak}$  per kg lean body mass has been suggested to be the best expression of aerobic capacity because it theoretically makes sense to express aerobic capacity relative to the maximal work capacity of muscle (not including fat tissue), and because this measure is empirically weaker as related to fatness than  $VO_{2peak}$  per kg body mass<sup>44</sup>. Alternatively, allometric scaling (raising body mass to a power function) can be used to reduce the relation to body size and fatness<sup>44</sup>.  $VO_{2peak}$  per kg body mass is generally negatively related to indices of fatness<sup>44 45</sup> ( $r = -0.23$  with body mass and  $r = -0.42$  with BMI in the present study after adjustment for age and sex, results not shown), meaning that  $VO_{2peak}$  per kg body mass in relation to metabolic health is confounded by fatness. Thus, studies in children have shown that associations with metabolic health indices have been attenuated when expressing  $VO_{2peak}$  per kg lean mass versus body mass or when controlling for fatness<sup>43 45 46</sup>. Contrary to these studies, which sought to determine the association between aerobic fitness and diverse indicators of metabolic risk, we present herein cut points that can be used to classify children at risk. Thus, the abovementioned studies and the present study are answering two different questions. One question is whether cardiorespiratory fitness is related to metabolic health *per se*. The other is whether aerobic fitness could predict metabolic risk and thus be used for screening in healthy populations. This is a fundamental distinction, because the first question is a question of etiology (where confounding should be removed), whereas the other is a question of prediction (where confounding is no problem or actually could be seen as a strength). If  $VO_{2peak}$  were 100% confounded by a variable strongly related to metabolic health,  $VO_{2peak}$  would be an excellent predictor and the relating cut points would consequently do an excellent job in discriminating between those at risk and those not at risk. The present study's research question is one of prediction; we have therefore not attempted to remove possible confounding as in this case it does not make sense to remove information relevant for the outcome.

## Conclusion

We present  $VO_{2peak}$  (ml/kg/min) cut points to classify children and youth with poor cardiometabolic health without using invasive measures of traditional CVD risk factors. Together with BMI, these cut points can be used to screen schoolchildren for poor cardiometabolic health and thereby inform teachers and health authorities who to target with physical activity initiatives. However, screening deserves thorough consideration of the target population and the approach for intervention, given the moderate discriminating performance for current cardiometabolic risk and uncertain future risk.

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